On the Double Descent of Random Features Models Trained with SGD

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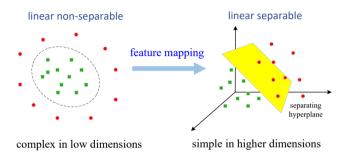
Outline

Research overview

Random features in double descen-

Conclusion

Research Overview: Kernel approximation



Scalability of kernel methods: n-by-n kernel matrix. Solution: approximate the kernel by a low-rank representation

- ▶ Nyström approximation: approximate the kernel matrix
- ▶ Random Fourier features¹: approximate the kernel function

¹Rahimi A, Recht B. Random features for large-scale kernel machines, NeurIPS2007. (the test-of-time award in NeurIPS2017)

Research Overview: Random Fourier features

$$k(\mathbf{x}, \mathbf{x}') = \langle \phi(\mathbf{x}), \phi(\mathbf{x}') \rangle_{\mathcal{H}} \approx \varphi^{\top}(\mathbf{x}) \varphi(\mathbf{x}'),$$

where $\varphi(\mathbf{x}): \mathbb{R}^d \to \mathbb{R}^s$ is an **explicit** feature mapping

Bochner's theorem [1]

For a shift-invariant $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x} - \mathbf{x}')$ and positive definite kernel,

$$k(\mathbf{x}, \mathbf{x}') = \int_{\mathbb{R}^d} \mathbf{p}(\boldsymbol{\omega}) \exp\left(i\boldsymbol{\omega}^\top (\mathbf{x} - \mathbf{x}')\right) d\boldsymbol{\omega}$$
$$\approx \frac{1}{s} \sum_{j=1}^s \exp(i\boldsymbol{\omega}_j^\top \mathbf{x}) \exp(i\boldsymbol{\omega}_j^\top \mathbf{x}')^* = \varphi(\mathbf{x})^\top \varphi(\mathbf{x}')$$

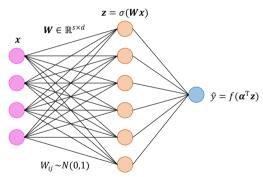
the explicit feature mapping:

$$\varphi(\mathbf{x}) := \frac{1}{\sqrt{s}} \left[\exp(-\mathsf{i}\omega_1^\top \mathbf{x}), \cdots, \exp(-\mathsf{i}\omega_s^\top \mathbf{x}) \right]^\top.$$

Research Overview: Neural network view

RF model: a two-layer, (infinite)-width, fully-connected neural network

$$k\left(\mathbf{x}, \mathbf{x}'\right) = \mathbb{E}_{\omega \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)}[\sigma(\omega^{\top} \mathbf{x}) \sigma(\omega^{\top} \mathbf{x}')]$$



- Gaussian kernel: $\sigma(x) = [\cos(x), \sin(x)]^{\top}$
- ▶ the 1st-order arc-cosine kernel: $\sigma(x) = \max\{0, x\}$
- ightharpoonup soft-max in attention: $\sigma(x) = \exp(x)$

Research Overview: Applied to Linearized Attention in Transformers

self attention

$$\operatorname{Attention}(\mathbf{Q}, \mathbf{K}, \mathbf{V}) = \underbrace{\operatorname{softmax}(\mathbf{Q}\mathbf{K}^\top)}_{:=\mathbf{A}} \mathbf{V} \approx \mathbf{Q}'\mathbf{K}'^\top\mathbf{V} \,,$$

where $\mathbf{A}_{ij} = k(\mathbf{q}_i, \mathbf{k}_j) = \mathbb{E}[\sigma(\mathbf{q}_i)^{\top} \sigma(\mathbf{k}_j)]$

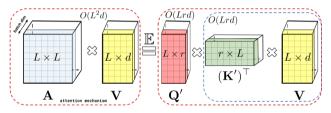


Figure: Approximation of self-attention. source: [2].

 $\qquad \qquad \mathbf{soft\text{-max in attention: }} \exp(\mathbf{x}^{\!\top}\mathbf{x}') = \mathbb{E}_{\omega \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_d)} \left[\exp\left(\omega^{\!\top}\mathbf{x} - \frac{\|\mathbf{x}\|_2^2}{2}\right) \exp\left(\omega^{\!\top}\mathbf{x}' - \frac{\|\mathbf{x}'\|_2^2}{2}\right) \right]$

Research Overview: Taxonomy

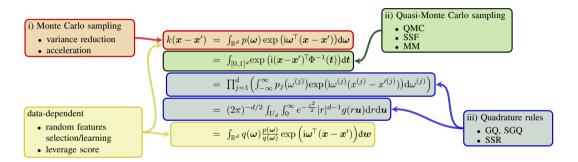


Figure: Taxonomy of random features based algorithms².

²Fanghui Liu, Xiaolin Huang, Yudong Chen, and Johan A.K. Suykens. *Random Features for Kernel Approximation: A Survey on Algorithms, Theory, and Beyond.* TPAMI2021.

Outline

Research overview

Random features in double descent

Conclusion

Background: Double descent

over-parameterized models, e.g., neural networks, random features

- ightharpoonup high dimensions: large n and d
- ▶ abnormal phenomena: training error can be zero but still generalize well

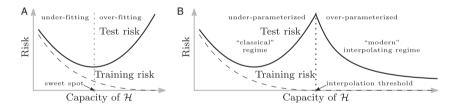


Figure: Bias-variance trade-off [3] (Belkin et al. PNAS2019).

Research Overview: Motivation

- interplay between optimization and excess risk: trained by SGD
- bias-variance decomposition for understanding multiple randomness sources

	data assumption	solution	result
(Hastie et al., 2019)	Gaussian	closed-form	variance 🗡 🔪
(Ba et al., 2020)	Gaussian	GD	variance 🗡 🔪
(Mei & Montanari, 2019)	i.i.d on sphere	closed-form	variance, bias 🗡 🤪
(d'Ascoli et al., 2020a)	Gaussian	closed-form	refined 2
(Gerace et al., 2020)	Gaussian	closed-form	7 \
(Adlam & Pennington, 2020)	Gaussian	closed-form	refined
(Dhifallah & Lu, 2020)	Gaussian	closed-form	7 😾
(Hu & Lu, 2020)	Gaussian	closed-form	7 \
(Liao et al., 2020)	general	closed-form	7 \
(Lin & Dobriban, 2021)	isotropic features with finite moments	closed form	refined
(Li et al., 2021)	correlated features with polynomial decay on Σ_d	closed form	interpolation learning
Ours	(at least) sub-exponential data	SGD	variance / , bias

A refined decomposition on variance is conducted by sources of randomness on data sampling, initialization, label noise to possess each term (d'Ascoli et al., 2020b) or their full decomposition in (Adlam & Pennington, 2020; Lin & Dobriban, 2021).

Problem settings: Random features regression model

data: $y = f_{\rho}(\mathbf{x}) + \varepsilon$

- ▶ training data: $\{(\mathbf{x}_i, y_i)\}_{i=1}^n \sim \rho$ Assumption: sub-exponential data and $\|\mathbf{x}\|_2^2 \sim \mathcal{O}(d)$
- ▶ target function: $f_{\rho}(\mathbf{x}) = \int_{V} y \, d\rho(y \mid \mathbf{x})$
- ▶ noise: $\mathbb{E}(\varepsilon) = 0$ and $\mathbb{E}(\varepsilon^2) = \tau^2$

function space

define the random features mapping $\varphi(\mathbf{x}) := \frac{1}{\sqrt{m}} \sigma(\mathbf{W}\mathbf{x}/\sqrt{d})$,

$$\mathcal{H} := \left\{ f \in L^2_{\rho_X} \left| \ f(\mathbf{x}) = \langle \theta, \varphi(\mathbf{x}) \rangle \right. \right\} \,, \quad \mathbf{W}_{ij} \sim \mathcal{N}(0, 1)$$

covariance operator: $\Sigma_m := \int_X [\varphi(\mathbf{x}) \otimes \varphi(\mathbf{x})] \mathrm{d}\rho_X(\mathbf{x})$

expected covariance operator: $\widehat{\Sigma}_m := \mathbb{E}_{\mathbf{x},\mathbf{W}}[\varphi(\mathbf{x}) \otimes \varphi(\mathbf{x})]$

Problem settings: averaged SGD under adaptive step-size setting

$$\theta_t = \theta_{t-1} + \gamma_t [y_t - \langle \theta_{t-1}, \varphi(\mathbf{x}_t) \rangle] \varphi(\mathbf{x}_t), \qquad t = 1, 2, \dots n,$$

- ▶ averaged output: $\bar{\theta}_n := \frac{1}{n} \sum_{t=0}^{n-1} \theta_t \Longrightarrow \bar{f}_n = \langle \varphi(\cdot), \bar{\theta}_n \rangle$
- ▶ adaptive step-size: $\gamma_t := \gamma_0 t^{-\zeta}, \zeta \in [0,1)$
- ightharpoonup optimal solution: $f^* = \arg\min_{f \in \mathcal{H}} \|f f_{\rho}\|_{L_{\rho_X}^2}^2$
- $\qquad \text{averaged excess risk: } \mathbb{E} \| \bar{f}_n f^* \|_{L^2_{\rho_X}}^2 = \mathbb{E}_{\mathbf{X}, \mathbf{W}, \varepsilon} \langle \bar{f}_n f^*, \Sigma_m(\bar{f}_n f^*) \rangle$

Assumptions

- **boundedness of** f^* : $||f^*||_{\mathcal{H}} < \infty$
- ▶ high dimension: $c \leq \{d/n, m/n\} \leq C$, $\|\mathbf{x}\|_2^2 \sim \mathcal{O}(d)$, $\Sigma_d := \mathbb{E}_{\mathbf{x}}[\mathbf{x} \otimes \mathbf{x}]$ with $\|\Sigma_d\|_2 < \infty$
- **activation function:** $\sigma(\cdot)$: Lipschitz continuous
- ▶ noise condition: $\Xi := \mathbb{E}_{\mathbf{x}}[\varepsilon^2 \varphi(\mathbf{x}) \otimes \varphi(\mathbf{x})] \leq \tau^2 \Sigma_m$. uniformly bounded noise, sub-Gaussian noise
- ► fourth moment condition:

for any PSD operator A, we have $\mathbb{E}_{\mathbf{W}}[\Sigma_m A \Sigma_m] \leq r' \mathbb{E}_{\mathbf{W}}[\mathrm{Tr}(\Sigma_m A) \Sigma_m] \leq r \mathrm{Tr}(\widetilde{\Sigma}_m A) \widetilde{\Sigma}_m$.

- 1) The special case A := I can be proved.
- 2) holds for sub-Gaussian/exponential data.

Properties of covariance operators

$$\begin{split} \sigma(\cdot) : \mathbb{R} &\mapsto \mathbb{R} \text{ Lipschitz continuous} \\ \text{covariance operator } \Sigma_m := \mathbb{E}_{\mathbf{x}}[\varphi(\mathbf{x}) \otimes \varphi(\mathbf{x})] \\ \text{expected covariance operator } \widetilde{\Sigma}_m := \mathbb{E}_{\mathbf{x},\mathbf{W}}[\varphi(\mathbf{x}) \otimes \varphi(\mathbf{x})] \end{split}$$

eigenvalue of $\widetilde{\Sigma}_m$

the same diagonal/non-diagonal elements: $\mathcal{O}(1/m)$

two distinct eigenvalues: $\widetilde{\lambda}_1 \sim \mathcal{O}(1), \ \widetilde{\lambda}_2 \sim \mathcal{O}(1/m)$

sub-exponential random variables

 $\|\Sigma_m\|_2$, $\|\Sigma_m - \widetilde{\Sigma}_m\|_2$, $\mathrm{Tr}(\Sigma_m)$, and $\left\|\widetilde{\Sigma}_m^{-1}\mathbb{E}_{\mathbf{W}}(\Sigma_m^2)\right\|_2$ with $\mathcal{O}(1)$ sub-exponential norm order

Bias-variance decomposition

Define $\eta_t := f_t - f^*$, we have

$$\begin{split} \eta_t &= [I - \gamma_t \varphi(\mathbf{x}_t) \otimes \varphi(\mathbf{x}_t)] (f_{t-1} - f^*) + \gamma_t \varepsilon_t \varphi(\mathbf{x}_t) \,, \\ \eta_t^{\texttt{bias}} &= [I - \gamma_t \varphi(\mathbf{x}_t) \otimes \varphi(\mathbf{x}_t)] \eta_{t-1}^{\texttt{bias}}, \quad \eta_0^{\texttt{bias}} = f^* \,, \\ \eta_t^{\texttt{var}} &= [I - \gamma_t \varphi(\mathbf{x}_t) \otimes \varphi(\mathbf{x}_t)] \eta_{t-1}^{\texttt{var}} + \gamma_t \varepsilon_t \varphi(\mathbf{x}_t), \quad \eta_0^{\texttt{var}} = 0 \,. \end{split}$$

Bias-variance decomposition

$$\mathbb{E} \|\bar{f}_n - f^*\|_{L^2_{\rho_X}}^2 = \underbrace{\mathbb{E}_{\mathbf{X}, \mathbf{W}} \langle \bar{\eta}_n^{\mathtt{bias}}, \Sigma_m \bar{\eta}_n^{\mathtt{bias}} \rangle}_{:=\mathtt{Bias}} + \underbrace{\mathbb{E}_{\mathbf{X}, \mathbf{W}, \varepsilon} \langle \bar{\eta}_n^{\mathtt{var}}, \Sigma_m \bar{\eta}_n^{\mathtt{var}} \rangle}_{:=\mathtt{Variance}}.$$

Proof framework

$$\begin{array}{c} \left(\begin{array}{c} \text{excess risk } \mathbb{E}_{\boldsymbol{X},\boldsymbol{W},\boldsymbol{\varepsilon}} \langle \bar{\eta}_{n}, \boldsymbol{\Sigma}_{m} \bar{\eta}_{n} \rangle \end{array} \right) \\ & \left(\begin{array}{c} \text{Variance } \mathbb{E}_{\boldsymbol{X},\boldsymbol{W},\boldsymbol{\varepsilon}} \langle \bar{\eta}_{n}^{\text{var}}, \boldsymbol{\Sigma}_{m} \bar{\eta}_{n}^{\text{var}} \rangle \end{array} \right) \\ & \left(\begin{array}{c} \text{V1: } \bar{\eta}_{n}^{\text{var}} - \bar{\eta}_{n}^{\text{vx}} \\ \mathcal{O}(n^{\zeta-1}m) & \text{if } m \leqslant n \\ \mathcal{O}(1) & \text{if } m \geqslant n \end{array} \right) \\ \left(\begin{array}{c} \text{V2: } \bar{\eta}_{n}^{\text{vx}} - \bar{\eta}_{n}^{\text{vx}} \\ \mathcal{O}(n^{\zeta-1}m) \\ \mathcal{O}(n^{\zeta-1} + \frac{n}{m}) \end{array} \right) \\ \left(\begin{array}{c} \text{V3: } \bar{\eta}_{n}^{\text{vxw}} - \bar{\eta}_{n}^{\text{bx}} \\ \mathcal{O}(n^{\zeta-1}m) \\ \mathcal{O}(n^{\zeta-1}m) \\ \mathcal{O}(n^{\zeta-1}) \end{array} \right) \\ \left(\begin{array}{c} \text{B1: } \bar{\eta}_{n}^{\text{bias}} - \bar{\eta}_{n}^{\text{bx}} \\ \mathcal{O}(n^{\zeta-1}m) \\ \mathcal{O}(n^{\zeta-1}) \end{array} \right) \\ \left(\begin{array}{c} \text{B3: } \bar{\eta}_{n}^{\text{bxy}} \\ \mathcal{O}(n^{\zeta-1}) \\ \mathcal{O}(n^{\zeta-1}) \end{array} \right) \\ \left(\begin{array}{c} \text{B3: } \bar{\eta}_{n}^{\text{bxy}} \\ \mathcal{O}(n^{\zeta-1}) \\ \mathcal{O}(n^{\zeta-1}) \end{array} \right) \\ \left(\begin{array}{c} \text{Constant } \\ \mathcal{O}(n^{\zeta-1}m) \\ \mathcal{O}(n^{\zeta-1}) \end{array} \right) \\ \left(\begin{array}{c} \text{Constant } \\ \mathcal{O}(n^{\zeta-1}) \\ \mathcal{O}(n^{\zeta-1}) \end{array} \right) \\ \left(\begin{array}{c} \text{Constant } \\ \mathcal{O}(n^{\zeta-1}) \\ \mathcal{O}(n^{\zeta-1}) \end{array} \right) \\ \left(\begin{array}{c} \text{Constant } \\ \mathcal{O}(n^{\zeta-1}) \\ \mathcal{O}(n^{\zeta-1}) \\ \mathcal{O}(n^{\zeta-1}) \end{array} \right) \\ \left(\begin{array}{c} \text{Constant } \\ \mathcal{O}(n^{\zeta-1}) \\ \mathcal{O}(n^{$$

$$\text{Bias}: \quad \eta_t^{\text{bias}} = [I - \gamma_t \varphi(\mathbf{x}_t) \otimes \varphi(\mathbf{x}_t)] \eta_{t-1}^{\text{bias}}$$

Define "semi-stochastic" version: $\eta_t^{\mathrm{bX}} = (I - \gamma_t \Sigma_{\mathbf{m}}) \eta_{t-1}^{\mathrm{bX}}, \quad \eta_t^{\mathrm{bXW}} = (I - \gamma_t \widetilde{\Sigma}_{\mathbf{m}}) \eta_{t-1}^{\mathrm{bXW}},$

- $\blacktriangleright \ \mathtt{B1} := \mathbb{E}_{\mathbf{X},\mathbf{W}} \Big[\langle \bar{\eta}_n^{\mathtt{bias}} \bar{\eta}_n^{\mathtt{bX}}, \Sigma_m (\bar{\eta}_n^{\mathtt{bias}} \bar{\eta}_n^{\mathtt{bX}}) \rangle \Big]$
- $\blacktriangleright \ \mathtt{B2} := \mathbb{E}_{\mathbf{W}} \left[\langle \bar{\eta}_n^{\mathtt{bX}} \! \! \bar{\eta}_n^{\mathtt{bXW}}, \Sigma_m (\bar{\eta}_n^{\mathtt{bX}} \! \! \bar{\eta}_n^{\mathtt{bXW}}) \rangle \right]$
- \triangleright B3 := $\langle \bar{\eta}_n^{\text{bXW}}, \widetilde{\Sigma}_m \bar{\eta}_n^{\text{bXW}} \rangle$

Proof framework

$$\begin{array}{c} \left(\begin{array}{c} \text{Excess risk } \mathbb{E}_{\boldsymbol{X},\boldsymbol{W},\boldsymbol{\varepsilon}} \langle \bar{\eta}_n, \boldsymbol{\Sigma}_m \bar{\eta}_n \rangle \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Variance } \mathbb{E}_{\boldsymbol{X},\boldsymbol{W},\boldsymbol{\varepsilon}} \langle \bar{\eta}_n^{\text{var}}, \boldsymbol{\Sigma}_m \bar{\eta}_n^{\text{var}} \rangle \\ \end{array} \right) \\ \left(\begin{array}{c} \text{V1: } \bar{\eta}_n^{\text{var}} - \bar{\eta}_n^{\text{vX}} \\ \mathcal{O}(n^{\zeta-1}m) & \text{if } m \leqslant n \\ \mathcal{O}(1) & \text{if } m \geqslant n \end{array} \right) \\ \left(\begin{array}{c} \text{V2: } \bar{\eta}_n^{\text{vX}} - \bar{\eta}_n^{\text{vXW}} \\ \mathcal{O}(n^{\zeta-1}m) \\ \mathcal{O}(n^{\zeta-1} + \frac{n}{m}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{V3: } \bar{\eta}_n^{\text{vXW}} - \bar{\eta}_n^{\text{bias}} - \bar{\eta}_n^{\text{bX}} \\ \mathcal{O}(n^{\zeta-1}m) \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{B3: } \bar{\eta}_n^{\text{bXW}} - \bar{\eta}_n^{\text{bXW}} \\ \mathcal{O}(n^{\zeta-1}m) \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{B3: } \bar{\eta}_n^{\text{bXW}} - \bar{\eta}_n^{\text{bXW}} \\ \mathcal{O}(n^{\zeta-1}) \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1}) \\ \end{array} \right) \\ \left(\begin{array}{c} \text{Column } \\ \mathcal{O}(n^{\zeta-1})$$

$$\text{Variance}: \quad \eta_t^{\text{var}} = [I - \gamma_t \varphi(\mathbf{x}_t) \otimes \varphi(\mathbf{x}_t)] \eta_{t-1}^{\text{var}} + \gamma_t \varepsilon_t \varphi(\mathbf{x}_t)$$

Define "semi-stochastic" version: $\eta_t^{\text{vX}} := (I - \gamma_t \underline{\Sigma_m}) \eta_{t-1}^{\text{vX}} + \gamma_t \varepsilon_t \varphi(\mathbf{x}_t), \quad \eta_t^{\text{vXW}} := (I - \gamma_t \underline{\widetilde{\Sigma}_m}) \eta_{t-1}^{\text{vXW}} + \gamma_t \varepsilon_t \varphi(\mathbf{x}_t)$

- $\blacktriangleright \ \mathtt{V1} := \mathbb{E}_{\mathbf{X},\mathbf{W},\varepsilon} \left[\langle \bar{\eta}_n^{\mathtt{var}} \bar{\eta}_n^{\mathtt{vX}}, \Sigma_m (\bar{\eta}_n^{\mathtt{var}} \bar{\eta}_n^{\mathtt{vX}}) \rangle \right]$
- $ightharpoonup V3 := \mathbb{E}_{\mathbf{X}} \mathbf{W} \in \langle \bar{\eta}_n^{\text{vXW}}, \Sigma_m \bar{\eta}_n^{\text{vXW}} \rangle$

Results: error bounds

Theorem

Under the above-mentioned assumptions, if the step-size $\gamma_t := \gamma_0 t^{-\zeta}$ with $\zeta \in [0,1)$ satisfies $\gamma_0 < C$, we have

$$\begin{split} \operatorname{Bias} &\lesssim \frac{\gamma_0 r' n^{\zeta-1}}{\sqrt{\mathbb{E}[1-\gamma_0 r' \operatorname{Tr}(\Sigma_m)]^4}} \|f^*\|^2 \sim \mathcal{O}\left(n^{\zeta-1}\right) \,. \\ \operatorname{Variance} &\lesssim \frac{\gamma_0 r' \tau^2}{\sqrt{\mathbb{E}[1-\gamma_0 r' \operatorname{Tr}(\Sigma_m)]^2}} \left\{ \begin{array}{l} m n^{\zeta-1}, \ \text{if} \ m \leqslant n \\ \gamma_0 \tau^2, \ \text{if} \ m > n \end{array} \right. \\ &\sim \left\{ \begin{array}{l} \mathcal{O}\left(m n^{\zeta-1}\right), \ \text{if} \ m \leqslant n \\ \mathcal{O}(1), \ \text{if} \ m > n \end{array} \right. \end{split}$$

Experiments on MNIST

Gaussian kernel $k(\mathbf{x}, \mathbf{x}') = \exp\left(-\frac{\|\mathbf{x} - \mathbf{x}'\|_2^2}{2d}\right)$

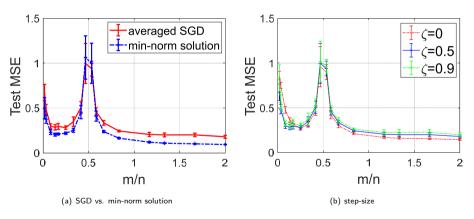
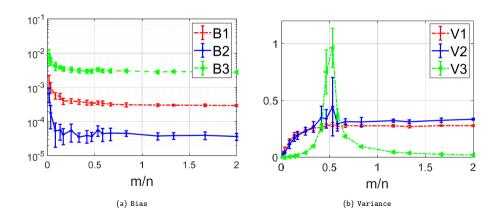


Figure: Test MSE (mean \pm std.) of RF regression as a function of the ratio m/n on MNIST data set (digit 3 vs. 7) for d=784 and n=600.

Validation for bias and variance

- ▶ noise: $\varepsilon \sim \mathcal{N}(0,1)$
- $ightharpoonup \Sigma_m$: sample covariance matrices with Monte Carlo sampling



Outline

Conclusion

Take-away message

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\begin{cases} \text{high dimensional random features model trained by SGD} \\ \\ \text{findings} \\ \end{cases} \begin{cases} \text{expected covariance operator } \widetilde{\Sigma}_m \text{ has only two distinct eigenvalues} \\ \\ \text{bias-variance decomposition: multiple randomness sources} \\ \\ \text{monotonic decreasing bias and unimodal variance} \\ \\ \text{optimization effect on excess risk: constant step-size SGD vs. min-norm solution} \end{cases}
```

Future works:

- ► SGD: implicit bias/regularization
- function space, high dimensions

Thanks for your attention!

Q & A

my homepage http://lfhsgre.org for more information!



NEW: ERC Advanced Grant E-DUALITY Exploring duality for future data-driven modelling













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